

DERIVATION AND PROPERTIES OF THE SIZE-BIASED POISSON JANARDAN

DISTRIBUTION

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ABSTRACT

In this paper the size-biased Poisson Janardan distribution (SBPJD) is introduced. The probability distribution of size-biased Poisson Janardan distribution is obtained by considering size-biased form of the Poisson distribution and Janardan distribution without its size-biased form. Some of its basic properties are derived and it is found that size-biased Poisson Lindley distribution given by Srivastava and and Adhikari, is a special case of the size-biased Poisson Janardan distribution. The equations of the method of moment and maximum likelihood estimators are obtained to find the estimators of the parameters of the size-biased Poisson Janardan distribution.

KEYWORDS: MLE, MOM, Poisson Distribution, Size-Biased Distribution, SBPJD, SBPLD

1. INTRODUCTION

Shanker (2013) obtained the two parameter Janardan distribution that is the mixture of exponential $\left(\frac{\theta}{\alpha}\right)$ and

gamma $\left(2, \frac{\theta}{\alpha}\right)$ distributions with the following probability distribution function (pdf)

$$f(x;\theta,\alpha) = \frac{\theta^2}{\alpha(\theta+\alpha^2)} (1+\alpha x) e^{-\frac{\theta}{\alpha}x}, \quad x > 0, \quad \theta > 0, \alpha > 0.$$
(1.1)

Lindley (1958) introduced a one parameter distribution named Lindley distribution with pdf

$$f(x;\theta) = \frac{\theta^2}{\theta+1} (1+x)e^{-\theta x}, \quad x > 0, \quad \theta > 0.$$

$$(1.2)$$

It can be seen that for $\alpha = 1$, the Janardan distribution (1.1) is becomes Lindley distribution (1.2).

Shanker et al (2014) obtained the discrete Poisson Janardan distribution (PJD) as

$$f(x;\theta,\alpha) = \left(\frac{\theta}{\theta+\alpha}\right)^2 \left(\frac{\alpha}{\theta+\alpha}\right)^x \left(1 + \frac{\alpha(1+\alpha x)}{\theta+\alpha^2}\right), \quad x = 0,1,2,\dots,; \quad \theta > 0, \alpha > 0.$$
(1.3)

For $\alpha = 1$, the PJD in (1.3) is becomes the Poisson Lindley distribution (PLD) introduced by Sankaran (1970).

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The pdf of size-biased Poisson distribution (SBPD) is

$$P(x;\lambda) = e^{-\lambda} \frac{\lambda^{x-1}}{(x-1)!}, \qquad x = 1,2,3,...,; \quad \lambda > 0.$$
(1.4)

Adhikari and Srivastava (2013) introduced the size-biased Poisson Lindley distribution (SBPLD) which is obtained by considering the size-biased Poisson distribution with Lindley distribution without considering its size-biased form. Then the pdf of SBPLD

$$f(x;\theta) = \frac{\theta^2}{(1+\theta)^{x+2}} (x+\theta+1), \qquad x = 1,2,3....,; \quad \theta > 0.$$
(1.5)

Adhikari and Srivastava (2014) proposed Poisson size-biased Lindley distribution (PSBLD) considering Poisson without its size-biased version and size-biased version of Lindley. Ghitany and Al-Mutairi (2008) proposed the size-biased Poisson Lindley distribution which is obtained by considering the size-biased version of Poisson Lindley distribution. They discussed the estimation methods for the size-biased Poisson Lindley distribution and its applications on real data sets.

Suppose that the original values of x comes from a distribution with pdf $f_0(x)$ and the values of x recorded according to a probability re-weighted by a weight function w(x) > 0, then the pdf of x

$$f(x) = \frac{w(x)}{E(w(x))} f_0(x)$$

Rao (1965) introduced this type of distributions and named them weighted distributions. For w(x) = x, is called size-biased or length-biased distribution. Patil and Rao (1978) showed that the size-biased distributions occur in a usual way in many sampling problems. Patil and Ord (1975) discussed size-biased and related weighted distributions in sampling procedures. Patil and Rao (1977) discussed the applications of size-biased distributions in real-life problems.

2. SIZE-BIASED POISSON JANARDAN DISTRIBUTION (SBPJD)

Suppose that λ of the size biased Poisson distribution in (1.4) follows the Janardan distribution in (1.1). Then the mixture of size-biased Poisson and two parameter Janardan distributions is obtained as

$$f(x;\theta,\alpha) = \int_{0}^{\infty} P(x;\lambda) f(\lambda;\theta,\alpha) d\lambda.$$

$$f(x;\theta,\alpha) = \frac{\theta^{2}}{\alpha(\theta+\alpha^{2})} \left(\frac{\alpha}{\theta+\alpha}\right)^{x} \left(1 + \frac{\alpha^{2}x}{\theta+\alpha}\right), \quad x = 1,2,3,...,; \quad \theta > 0, \alpha > 0.$$
(2.1)

The size-biased Poisson Janardan distribution (SBPJD) in (2.1) is reduces to size-biased Poisson Lindley distribution (SBPLD) in (1.5).

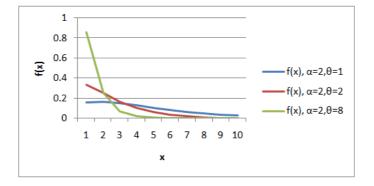


Figure 1: Plot of Pdf SBPLD For $\alpha = 2 \& \theta = 1,2,8$.

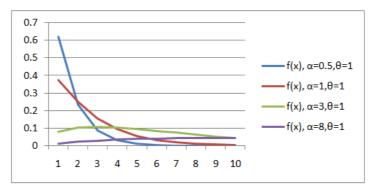


Figure 2: Plot of pdf SBPLD For $\alpha = 0.5, 1, 3, 8 \& \theta = 1$.

From Figure 1 & 2 it can be seen that the SBPJD is positively skewed. Figure 1 shows that for fix value of $\alpha = 2$ and $\theta = 8$ the pdf has high peaked with longer right tail and for $\theta = 1 \& 2$ the pdf is going to be flatter. Figure 2 shows that for fix value of $\theta = 1 \& \alpha = 0.5$ the pdf is high peaked with longer right tail and for $\alpha = 3 \& 8$ the pdf is going to be flatter. Moreover for $\alpha = 1$ the pdf is SBPLD.

Measures	SBPJD	SBPLD (For $\alpha = 1$ In SBPJD)
μ'_1	$1 + \frac{\alpha \left(\theta + 2\alpha^2\right)}{\theta \left(\theta + \alpha^2\right)}$	$\frac{\theta^2+2\theta+2}{\theta(\theta+1)}$
μ_2'	$\frac{\theta^2(\theta+\alpha^2)+3\theta\alpha(\theta+2\alpha^2)+2\alpha^2(\theta+3\alpha^2)}{\theta^2(\theta+\alpha^2)}$	$\frac{\theta^3 + 4\theta^2 + 8\theta + 6}{\theta^2(\theta+1)}$
μ'_3	$\frac{\theta^{3}(\theta + \alpha^{2}) + 7\theta^{2}\alpha(\theta + 2\alpha^{2}) + 12\theta\alpha^{2}(\theta + 3\alpha^{2}) + 6\alpha^{3}(\theta + 4\alpha^{2})}{\theta^{3}(\theta + \alpha^{2})}$	$\frac{\theta^4 + 8\theta^3 + 26\theta^2 + 42\theta + 24}{\theta^3(\theta+1)}$
μ'_4	$\frac{\theta^{4}(\theta + \alpha^{2}) + 15\theta^{3}\alpha(\theta + 2\alpha^{2}) + 50\theta^{2}\alpha^{2}(\theta + 3\alpha^{2}) + 60\theta\alpha^{3}(\theta + 4\alpha^{2}) + 24\alpha^{4}(\theta + 5\alpha^{2})}{\theta^{4}(\theta + \alpha^{2})}$	$\frac{\theta^5 + 16\theta^4 + 80\theta^3 + 210\theta^2 +}{264\theta + 120}$ $\frac{\theta^4(\theta + 1)}{\theta^4(\theta + 1)}$
σ^2	$\frac{\theta^{3}\alpha + \theta^{2}\alpha^{2} + 3\theta^{2}\alpha^{3} + 4\theta\alpha^{4} + 2\theta\alpha^{5} + 2\alpha^{6}}{\theta^{2}(\theta + \alpha^{2})^{2}}$	$\frac{\theta^3 + 4\theta^2 + 6\theta + 2}{\theta^2(\theta+1)^2}$

Table 1: The Moments of the SBPJD and SBPLD

Table 1 shows the first four moments, mean, variance and of the SBPJD and SBPLD

Some more properties of the size-biased Poisson Janardan distribution are

i. Since

$$\mu - \sigma^{2} = \frac{\theta^{4} + 2\theta^{3}\alpha^{2} + \theta^{2}\alpha^{4} - \theta^{2}\alpha^{2} - 4\theta\alpha^{4} - 2\alpha^{6}}{\theta^{2}(\theta + \alpha^{2})^{2}}$$
(2.2)

It follows from (2.2) that $\mu \ll \sigma^2$ for different values of $\alpha \& \theta$, that means SBPJD is over-dispersed/equidispersed/under-dispersed for different values of $\alpha \& \theta$.

	$\mu - \sigma^2$,	$\mu - \sigma^2$		
I		For $\theta = 1$		
$\theta = 0.05$	-798.0930	$\alpha = 0$		1
$\theta = 0.5$	-6.5556	$\alpha = 0.0001$		-0.9999
$\theta = 1$	-0.7500	$\alpha = 0.001$		-0.9999
$\theta = 1.5$	0.4311	$\alpha = 0.005$		-0.9999
$\theta = 2$	0.6111	$\alpha = 0.5$		-0.6
$\theta = 8$	0.9811	$\alpha = 2$		-1.8
$\theta = 10$	0.9883	$\alpha = 10$		-2.93
$\theta = 15$	0.9950			
$\theta = 20$	0.9973			
$\theta = 100$	0.99999 ≅ 1			

Table 2: Dispersion of the SBPJD

From the Table 2 it is observed that for fix value of $\alpha = 1$, the SBPJD is over-dispersed as the value of θ is increasing but the SBPJD is under-dispersed as the value of θ is decreasing. For fix value of $\theta = 1$, as the value of α increases the SBPJD is under-dispersed and the amount of dispersion increase as well. Therefore it can be seen that for $\theta = 100$, $\alpha = 1$ and $\theta = 1$, $\alpha = 0$, the SBPJD is qui-dispersed.

ii. Since

$$\frac{f(x+1;\theta,\alpha)}{f(x;\theta,\alpha)} = \frac{\alpha}{\theta+\alpha} \left(1 + \frac{\alpha^2}{\theta+\alpha+\alpha^2 x} \right)$$
(2.3)

It can be seen that eq(2.3) is a decreasing function in x, $f(x; \theta)$ is log concave. Therefore the SBPJD is unimodal.

The mode of the SBPJD is

$$\operatorname{mod} e = -\frac{1}{\ln\left(\frac{\alpha}{\theta + \alpha}\right)} - \frac{\theta + \alpha}{\alpha^{2}}$$
(2.4)

For $\alpha = 1$ it becomes the mode of the SBPLD.

3. METHOD OF MOMENTS (MOM)

Let a random sample of size n from SBPJD with pdf (2.1), the MOM estimates of the $\theta \& \alpha$ are respectively

$$\widetilde{\theta} = \frac{-\left[(\overline{x}-1)\alpha^2 - \alpha\right] + \sqrt{\left[(\overline{x}-1)\alpha^2 - \alpha\right] + 8\alpha^3(\overline{x}-1)}}{2(\overline{x}-1)}, \qquad \overline{x} > 1.$$
(2.5)

$$6\alpha^4 + 6\theta\alpha^3 + \theta(2 - \theta b + \theta)\alpha^2 + 3\theta^3\alpha - \theta^3(b - 1) = 0$$
(2.6)

$$b = \frac{\sum_{i=1}^{n} x_i^2}{2}$$

Where

4. METHOD OF MAXIMUM LIKELIHOOD ESTIMATOR (MLE)

Let a random sample of size n from SBPJD with pdf (2.1), the ML estimates of the $\theta \& \alpha$ are respectively

n

$$\frac{2n}{\theta} - \frac{n}{\theta + \alpha^2} - \frac{\sum_{i=1}^{n} x_i}{\theta + \alpha} - \frac{n}{\theta + \alpha} + \sum_{i=1}^{n} \frac{1}{\theta + \alpha + \alpha^2 x_i} = 0$$

$$\frac{n}{\alpha} + \frac{2n\alpha}{\theta + \alpha^2} - \frac{\sum_{i=1}^{n} x_i}{\alpha} + \frac{\sum_{i=1}^{n} x_i}{\theta + \alpha} + \frac{n}{\theta + \alpha} - \sum_{i=1}^{n} \frac{1 + 2\alpha x_i}{\theta + \alpha + \alpha^2 x_i} = 0$$
(2.7)
$$(2.7)$$

The ML estimate of $\hat{\theta}$ and $\hat{\alpha}$ are obtained by the solution of the non-linear equations

REFERENCES

- Ghitany, M. E. and Al-Mutairi, D. (2008). Size-biased Poisson Lindley distribution and its applications. Metron, LXVI, n. 3, 299-311.
- 2. Lindley, D, V. (1958). Fiducial distributions and Bayes theorem. Journal of Royal Statistical Society, 20, 102-107.
- 3. Patil, G. P. and Ord, J. K. (1975). On size-biased sampling and related form- invariant weighted distributions, Sankhya, 38, 48-61.
- 4. Patil, G. P. and Rao, C. R. (1977). Weighted distributions: a survey of their applications, in: Krishnaiah, P.R. (Ed), Applications of Statistics, Amsterdam, North-Holland, 383-405.
- 5. Patil, G. P. and Rao, C. R. (1978). Weighted distributions and size-biased sampling with applications to wildlife populations and human families, Biometrics, 34, 179-189.
- 6. Rao, C. R. (1965). On discrete distributions arising out of ascertainment, In: Classical and Contagious Discrete Distributions, Patil, G. P, (ed.), Pergamon Press and Statistical Publishing Society, Calcutta, 320-332.

- 7. Sankaran, M. (1970). The discrete Poisson Lindley distribution, Biometrics, 26, 145-149.
- 8. Shanker, R. and Fesshaye, H. (2015). On Poisson Lindley distribution and its applications to Biological sciences. Biometrics and Biostatistics International Journal, 2(4), 1-5.
- 9. Shanker, R. et al (2013). Janardan distribution and its applications to waiting time data. Indian Journal of Applied Research, 3(8), 500-502.
- Shanker, R. et al (2014). The discrete Poisson Janardan distribution with applications. International Journal of Soft Computing and Engineering, 4(2), 31-33.
- 11. Srivastava, R. S. and Adhikari, T. R. (2013). A size-biased Poisson Lindley distribution. (Accepted for International Journal of Multidisciplinary)
- Srivastava, R. S. and Adhikari, T. R. (2014). Poisson size-biased Lindley distribution. International Journal of Science and Research Publications, 4(1), 1-6.